

Exam Solutions

1. (Birthday problem)

Obviously $P(2 \text{ have same birthday}) = 1 - P(\text{all have different birthdays})$

$$\begin{aligned} P(\text{all have different}) &= \left(\frac{365}{365}\right) \left(\frac{364}{365}\right) \left(\frac{363}{365}\right) \dots \left(\frac{365-27+1}{365}\right) \\ &= \frac{1}{365^{27}} \cdot \frac{365!}{(365-27)!} \end{aligned}$$

$$P(2 \text{ have same}) = 1 - \frac{1}{365^{27}} \cdot \frac{365!}{(365-27)!} = 0.6269$$

2. (Proving a function is a pdf)

Using the theorem from the notes, we know that we only need to verify the two properties of the function $g(x)$ in order for it to be a pdf.

(i) Since $f_X(x)$ is a pdf, then $f_X(x) \geq 0$ for all x . By assumption, $1 - F_X(x_0) > 0$ and thus $g_X(x) \geq 0$ for all x .

$$(ii) \quad \int_{-\infty}^{\infty} g_X(x) dx = \frac{\int_{x_0}^{\infty} f_X(x) dx}{1 - F_X(x_0)} = \frac{1 - F_X(x_0)}{1 - F_X(x_0)} = 1$$

3. (Normal, chi-squared relationship)

$$EX^2 = \int_{-\infty}^{\infty} \left(\frac{1}{\sqrt{2\pi}}\right) e^{-x^2/2} x^2 dx = \left(\frac{1}{\sqrt{2\pi}}\right) \int_{-\infty}^{\infty} e^{-x^2/2} x^2 dx$$

We now must use change of variables and define $u = x^2$. Then notice $du = 2x dx$

$$\begin{aligned} EX^2 &= 2 \times \left(\frac{1}{\sqrt{2\pi}}\right) \int_0^{\infty} e^{-u/2} \frac{u}{2x} du \quad (\text{by symmetry}) \\ &= 2 \times \left(\frac{1}{\sqrt{2\pi}}\right) \frac{1}{2} \int_0^{\infty} e^{-u/2} (\sqrt{u}) du \end{aligned}$$

The easier method is to recall (or re-derive) $f_Y(y) = \frac{1}{\sqrt{2\pi}} \frac{1}{\sqrt{y}} e^{-y/2} \quad 0 < y < \infty$

$$\text{Then } EY = \int_0^{\infty} \frac{1}{\sqrt{2\pi}} \frac{y}{\sqrt{y}} e^{-y/2} dy$$

Note: I do not expect you to evaluate the above integral except through use of a table or a known fact. Notice that both EX^2 and EY yield the same value, in this case it is equal to 1. The fact about the Gamma function should come in handy here.

Define the change of variable $z = y/2$, then $dz = dy/2$.

$$\begin{aligned} EY &= \frac{1}{\sqrt{2\pi}} \times 2 \int_0^{\infty} (2z)^{1/2} e^{-z} dz = \frac{1}{\sqrt{\pi}} \times 2 \int_0^{\infty} z^{3/2-1} e^{-z} dz \\ &= \frac{1}{\sqrt{\pi}} \times 2 \times \Gamma\left(\frac{3}{2}\right) = \frac{2}{\sqrt{\pi}} \times \frac{1}{2} \times \Gamma\left(\frac{1}{2}\right) = \frac{2}{\sqrt{\pi}} \times \frac{\sqrt{\pi}}{2} = 1 \end{aligned}$$

4. (Uniform, exponential relationship)

$$EY = \int_0^1 g(x)dx = \int_0^1 -\log(x)dx$$

Now, we need to use integration by parts.

$$\text{Recall: } \int u dv = uv - \int v du$$

$$\text{Here, we set } \quad dv = -1 \quad u = \log(x)$$

$$\text{Then, of course } \quad v = -x \quad du = 1/x$$

$$\begin{aligned} \int_0^1 -\log(x)dx &= [-x \log(x)]_0^1 - \int_0^1 -dx \\ &= 0 + 0 + 1 = 1 \end{aligned}$$

Note: by L'Hospital's Rule $x \log(x)|_{x=0} = \frac{\log(x)}{1/x}|_{x=0}$;

$$\partial x \log(x)|_{x=0} = \frac{1/x}{-1/x^2}|_{x=0} = -x|_{x=0} = 0$$

5. (Moment generating functions)

$$\begin{aligned} M_X(t) &= \int_{-\infty}^{\infty} e^{xt} \frac{1}{\sqrt{2\pi}\sigma} e^{-(x-\mu)^2/(2\sigma^2)} dx = \frac{1}{\sqrt{2\pi}\sigma} \int_{-\infty}^{\infty} e^{(2\sigma^2 xt - x^2 + 2x\mu - \mu^2)/2\sigma^2} dx \\ &= \frac{1}{\sqrt{2\pi}\sigma} \int_{-\infty}^{\infty} e^{(-x^2 + 2x(\mu + \sigma^2 t) - (\mu + \sigma^2 t)^2)/2\sigma^2} dx \times e^{(2\mu\sigma^2 t + \sigma^4 t^2)/2\sigma^2} \\ &= 1 \times e^{\mu t + \sigma^2 t^2/2} \end{aligned}$$

6. (Application of Chebychev's Inequality)

Define $\epsilon = k\sigma$, then

$$P(|X - 3| \geq \epsilon) \leq \frac{\sigma^2}{\epsilon^2} \quad \forall \epsilon > 0$$

$$\text{know that } \sigma^2 = E(X^2) - E(X)^2 = 13 - 9 = 4$$

set $\epsilon = 5$, then

$$P(|X - 3| \geq 5) \leq \frac{4}{25}$$

$$1 - P(-2 < X < 8) \leq \frac{4}{25}$$

$$\Rightarrow P(-2 < X < 8) \geq \frac{21}{25}$$

7. (Find the MLE)

$$f_X(x; \alpha) = \frac{\partial F_X(x; \alpha)}{\partial x} = -\alpha \left(\frac{1}{x}\right)^{\alpha-1} \left(-\frac{1}{x^2}\right) = \alpha \left(\frac{1}{x}\right)^{\alpha+1}$$

$$L(\theta; X) = \prod_{i=1}^n \left(\alpha \left(\frac{1}{x_i}\right)^{\alpha+1} \right)$$

$$\begin{aligned} \log L(\theta; X) &= \sum_{i=1}^n \log(\alpha) - (\alpha + 1) \log(x_i) \\ &= n \log(\alpha) - (\alpha + 1) \sum_{i=1}^n \log(x_i) \end{aligned}$$

$$\frac{\partial \log L(\theta; X)}{\partial \alpha} = 0 \quad \Rightarrow \quad \frac{n}{\alpha} - \sum_{i=1}^n \log(x_i) = 0$$

$$\alpha^* = \frac{1}{\frac{1}{n} \sum_{i=1}^n \log(x_i)}$$